### Uncertainty, Real Activity, and Risk Aversion during the Great Recession

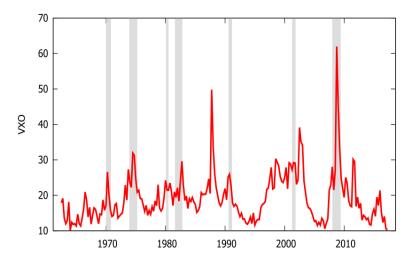
Giovanni Pellegrino Aarhus University

Giovanni Caggiano Monash University University of Padova

Efrem Castelnuovo University of Melbourne University of Padova

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#### Financial uncertainty and the business cycle



 Uncertainty during the Great Recession: Concern for policymakers Blanchard

#### Financial uncertainty and the Great Recession

- "But while policymakers clearly think uncertainty has played a central role in driving the Great Recession and slow recovery, the econometric evidence is really no more than suggestive." (Bloom, 2014)
- "An assumption of linearity may be adequate for estimating average relationships, but few expect that an economy will respond linearly to every aberration." (Greenspan, 29 August 2003)

#### Research questions

- ▶ Q1: Real effects of uncertainty shocks: State-dependent? Great Recession vs. normal times?
  - identification of new facts regarding the real effects of uncertainty shocks during the Great Recession
  - focus on financial uncertainty shocks, drivers of the business cycle (Ludvigson, Ma, and Ng 2019; Angelini, Bacchiocchi, Caggiano, and Fanelli 2019)
- ► Q2: Structural interpretation of asymmetric effects of uncertainty shocks during the Great Recession?
  - identification of structural drivers
  - calibration/estimation macro models for policy analysis

⇒Need nonlinear models

#### This paper

- Estimate IVARs with post-WWII U.S. data, GDP growth as conditioning indicator
  - ► Aim: State-dependent GIRFs to detect/quantify nonlinearities
- Estimate nonlinear DSGE model à la Basu and Bundick (2017) via Bayesian direct inference approach (Christiano et al., 2011)
  - Able to generate comovements to financial uncertainty shocks
  - Aim: Unveil whether/which structural instabilities in key parameters are needed to replicate the facts
- Analyze Great Recession
  - Output loss & implications for calibrating models for policy analysis, counterfactual on role of monetary policy

#### Main findings

- Responses of real activity indicators to uncertainty shocks: Stronger during Great Recession
  - Larger than predicted by a linear VAR
- ► Nonlinear effects of uncertainty shocks explained by counter-cyclical risk aversion
  - Good fit of the DSGE model in both states (normal times/GR) thanks to parameter instability
  - High risk aversion crucial to replicate IVAR Great Recession responses
- ► Great recession: Role of nonlinearities...
  - DSGE model estimated conditional on nonlinear IVAR explains about 60% of the output loss in the 2008Q4-2014Q3 period; only 40% if a linear VAR is employed
- ► Great recession: ... and policy counterfactual
  - ► Fed's immediate and aggresive response helped to prevent a deeper recession

#### Plan of the presentation

- Literature review
- ▶ IVAR model: Specification, estimation, evidence
- ▶ DSGE model: Review, estimation, evidence
- Great recession
- Conclusions

#### Literature review

- State-dependent effects of uncertainty shocks: Caggiano, Castelnuovo, and Groshenny (2014), Caggiano, Castelnuovo, Nodari (2017), Caggiano, Castelnuovo, Figueres (2017); Alessandri and Mumtaz (2018); Chatterjee (2018a)
- ► Interacted-VAR: Mittnik (1990); Towbin and Weber (2013), Sáet al. (2014), Aastveit et al. (2017), Pellegrino (2017a,b), Caggiano, Castelnuovo, Pellegrino (2017)
- ► Structural analysis of uncertainty-driven comovements: Basu and Bundick (2017)
- ► Bayesian IRF matching: Christiano et al. (2011); Castelnuovo and Pellegrino (2018)
- ► Estimation of III-order approx. DSGE frameworks: Fernández-Villaverde et al. (2011, 2015), Rudebusch and Swanson (2012), Born and Pfeifer (2014), Andreasen, Fernández-Villaverde, Rubio-Ramírez (2017), Ruge-Murcia (2017); Chatterjee (2018b); Mumtaz and Theodoridis (2019)

# Interacted VAR

#### Interacted VAR

► IVAR à la Pellegrino (2017a,b), Caggiano, Castelnuovo, and Pellegrino (2017):

$$\mathbf{Y}_t = \pmb{lpha} + \sum_{j=1}^L \mathbf{A}_j \mathbf{Y}_{t-j} + \left[ \sum_{j=1}^L \mathbf{c}_j \ln V X O_{t-j} imes \Delta \ln G D P_{t-j} 
ight] + \pmb{\eta}_t$$
 $E(\pmb{\eta}_t \pmb{\eta}_t') = \mathbf{\Omega}$ 

- $\mathbf{Y}_t = [\ln VXO_t, \ln GDP_t, \ln C_t, \ln I_t, \ln Hours_t, \ln P_t, R_t]'$
- Nonlinear effects of shocks to  $\ln VXO_t$  conditional on realizations of  $\Delta \ln GDP_{t-i}$
- Advantages of IVAR: Parsimonious framework, full sample estimation, resambles nonlineary decision rules of III-order model while preserving stability [results robust to IVAR with higher order terms]
- [Shadow rate consistent with a negative policy rate in DSGE model (Wu and Zhang 2017; Mouabbi and Sahuc 2017)]

#### Interacted-VAR (cont'd)

- Uncertainty can be (both) a cause and a consequence of business cycle fluctuations
- Uncertainty shocks identified using narrative sign restrictions as in Ludvigson et al. (LMN):

  - Infinitely many solutions:

$$\mathcal{B} = \left\{ \mathbf{B} = \mathbf{PQ} : \mathbf{Q} \in \mathcal{O}_{\underline{n}}, diag(\mathbf{B}) \geqslant 0, \mathbf{\Omega} = \mathbf{BB'} \right\}$$

$$ightharpoonup QQ' = I_n, \Omega = PP'$$

- ightharpoonup —Impose restrict. to get the set of admissible solutions  $ar{\mathcal{B}}$
- ▶ Constraints on  $\mathbf{e}_t(\mathbf{B})$ :
  - Event constraints: identified unc shock >75th (1987Q4, 2008Q4) [LMN] or 50th (Bloom's dates + 2016Q1) percentile of the empirical density in selected dates ( Dates )
  - External variable constraints: shocks must have negat. (posit.)
     correlation with stock market returns (gold price)

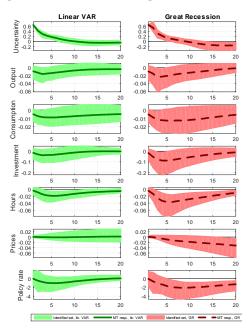
#### Interacted-VAR (cont'd)

▶ GIRFs as in Koop et al. (1996), Kilian and Vigfusson (2011):

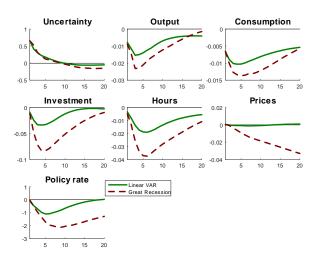
$$\textit{GIRF}_{\mathbf{y}}(h, \delta, \boldsymbol{\omega}_{t-1}) = E[\mathbf{y}_{t+h} | \delta, \boldsymbol{\omega}_{t-1}] - E[\mathbf{y}_{t+h} | \boldsymbol{\omega}_{t-1}]$$

- Focus on two responses:
  - great recession:  $GIRF_{y}(h, \delta, \omega_{2008Q3})$
  - nested linear VAR:  $IRF_{\mathbf{y}}(h, \delta)$
- $\delta = \delta_{2008Q4} = 4.4$  standard deviation uncertainty shock (• Distribution)
- ▶ Details: Sample: 1962Q3-2017Q4, VAR(4), OLS estimation
- ► Linearity rejected (pval < 0.01) via LR tests

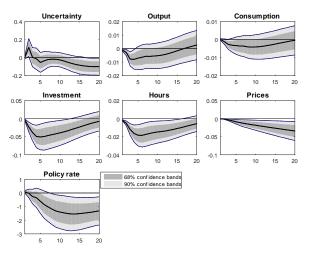
#### IVAR results: State-conditional GIRFs



#### IVAR results: Median Target GIRFs



#### IVAR results: Difference of state-conditional GIRFs



#### IVAR results: main results

- Real effects of financial uncertainty shocks stronger during the Great Recession
- Results robust to: credit spread, higher-order terms, money, fiscal stance, LMN financial uncertainty, time-varying VCV ( checks )

DSGE model-related interpretation?

## DSGE framework

#### DSGE model at a glance

- Medium-scale New Keynesian DSGE model à la Basu and Bundick (2017,2018)
  - sticky prices (Rotemberg), invest. adj. costs, var. cap. util.
- External habits in consumption added to generate hump-shaped responses
- EZ preferences:

$$\begin{array}{lll} V_t & = & \left[ (1-\beta) \left( a_t \tilde{C}_t^{\eta} (1-N_{t,\cdot})^{(1-\eta)} \right)^{\frac{1-\sigma}{\theta_V}} + \beta (E_t V_{t+1}^{1-\sigma})^{\frac{1}{\theta_V}} \right]^{\frac{\tau_V}{1-\sigma}} \\ \tilde{C}_t & = & C_t - H_t \text{ , with } H_t = bC_{t-1} \\ \theta_V & \equiv & (1-\sigma)/(1-\psi^{-1})^{-1} \end{array}$$

Demand-side uncertainty shock:

$$\begin{array}{lll} \mathbf{a}_t &=& (1-\rho_{\mathbf{a}})\mathbf{a} + \rho_{\mathbf{a}}\mathbf{a}_{t-1} + \sigma_{t-1}^{\mathbf{a}}\varepsilon_t^{\mathbf{a}} \\ \sigma_t^{\mathbf{a}} &=& (1-\rho_{\sigma^{\mathbf{a}}})\sigma^{\mathbf{a}} + \rho_{\sigma^{\mathbf{a}}}\sigma_{t-1}^{\mathbf{a}} + \sigma^{\sigma^{\mathbf{a}}}\varepsilon_t^{\sigma^{\mathbf{a}}} \end{array}$$

#### DSGE model at a glance (cont'd)

- ▶ Comovements due to an uncertainty shock, economic intuition:  $unc \uparrow, C \downarrow, L^s \uparrow, L^d =, w \downarrow, mc \downarrow, p \approx, markup \uparrow, Y \downarrow, L^d \downarrow, H \downarrow, I \downarrow$
- ▶ Model-implied VXO, match with our IVAR analysis:

$$VXO_{t}^{M}=100\cdot\sqrt{4\cdot VAR_{t}\left(R_{t+1}^{E}
ight)}$$

#### DSGE model: State-dependent estimation

- Christiano et al.'s (2011) Bayesian Minimum Distance approach applied to a III-order approximated model, state-dependent estimation
- IVAR GIRFs as "data". Under some assumptions, able to write the (limited information) likelihood and characterize the (quasi-) posterior density
- Our application allows for state-dependent parameters, employs an observable variable to identify the states, and is not computationally cumbersome
- ▶ Basu and Bundick (2017) nonlinear model estimated on the basis of state-dependent GIRFs

#### Christiano et al.'s (2011) Bayesian MD estimation

lacktriangle IVAR GIRFs  $\widehat{oldsymbol{\psi}^i}$  as "data", distribution as a function of the true DSGE model parameters  $\zeta_0$ :

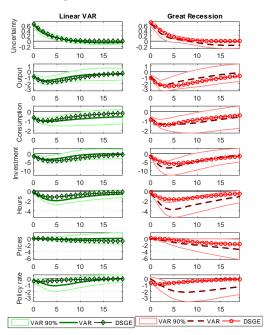
$$\widehat{oldsymbol{\psi}^i} \overset{ extstyle a}{\sim} extstyle N(oldsymbol{\psi}\left(oldsymbol{\zeta}_0^i
ight), oldsymbol{V}^i(oldsymbol{\zeta}_0^i, oldsymbol{n}^i)), ext{ for } i=R,E$$

$$\begin{split} f(\widehat{\boldsymbol{\psi}^{i}}|\boldsymbol{\zeta}^{i}) &= \left(\frac{1}{2\pi}\right)^{\frac{N^{i}}{2}} \left| \boldsymbol{\mathsf{V}}^{i}(\boldsymbol{\zeta}_{0}^{i},\boldsymbol{n}^{i}) \right|^{-\frac{1}{2}} \\ &\times \exp \left[ -\frac{1}{2} \left( \widehat{\boldsymbol{\psi}^{i}} - \boldsymbol{\psi} \left( \boldsymbol{\zeta}^{i} \right) \right)^{'} \boldsymbol{\mathsf{V}}^{i}(\boldsymbol{\zeta}_{0}^{i},\boldsymbol{n}^{i})^{-1} \left( \widehat{\boldsymbol{\psi}^{i}} - \boldsymbol{\psi} \left( \boldsymbol{\zeta}^{i} \right) \right) \right] \end{split}$$

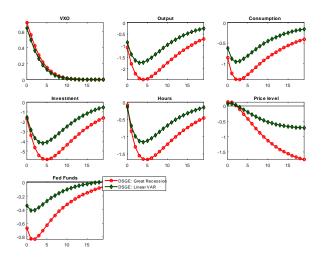
• 
$$\psi\left(\zeta^{i}
ight)$$
: DSGE-based GIRFs (III order perturbation)

$$p(\zeta^i) \text{ common between regimes, posterior density (MCMC):}$$
 
$$f(\zeta^i|\widehat{\psi^i}) = \frac{f(\widehat{\psi^i}|\zeta^i)p(\zeta^i)}{f(\widehat{\psi^i})}$$

#### DSGE fit: linear vs Great Recession



#### DSGE's GIRFs



#### DSGE model's estimated parameters

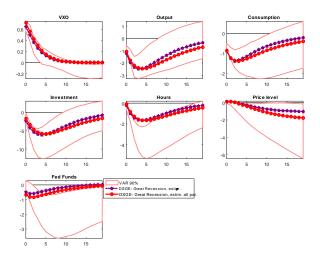
|                   |                   | Priors       | Posteriors           |                      |
|-------------------|-------------------|--------------|----------------------|----------------------|
|                   |                   |              | Linear VAR           | Great Recession      |
|                   |                   |              | All param.           | All param.           |
| Parameter         | Interpretation    | D(mean,std)  | Mode,std             | Mode,std             |
| $\rho_{\sigma^a}$ | Unc.shock,pers.   | B(0.77,0.10) | <b>0.64,</b> 0.03    | <b>0.65,</b> 0.03    |
| σ                 | Risk aversion     | G(100,60)    | <b>385.90,</b> 50.45 | <b>533.04,</b> 59.16 |
| Ь                 | Habits            | B(0.75,0.15) | <b>0.64,</b> 0.06    | <b>0.66,</b> 0.04    |
| $\phi_K$          | Inv. adj. costs   | G(3.92,2)    | <b>2.29,</b> 0.50    | <b>3.21,</b> 0.60    |
| $\phi_P$          | Price adj. costs  | G(240,40)    | <b>236.78,</b> 32.26 | <b>282.10,</b> 33.54 |
| $\rho_{\pi}$      | TR par., inflat.  | IG(1.5,0.25) | <b>1.05,</b> 0.01    | <b>1.05,</b> 0.01    |
| $\rho_{v}$        | TR par., out. gr. | G(0.2,0.15)  | <b>0.20,</b> 0.04    | <b>0.28,</b> 0.05    |

### DSGE results: Drivers

#### Role of state-dependent parameters

- ▶ Lots of instability in structural parameters, e.g. risk aversion, investment adj costs, price adj costs, CB response to output
- Relative role? Explorations conducted by swapping parameters one at at time point to risk aversion as crucial ( Counterfactual )
- ► Re-estimate our models by allowing for risk aversion only to be state-dependent; rest of the model calibrated with estimates conditional on linear VAR ( ► Distance criterion )

#### A risk aversion-only story?



#### DSGE model: Assessment

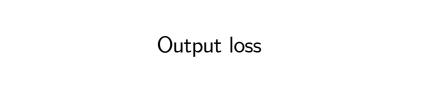
- Good performance of the model in replicating state-dependent IVARs
- ▶ Hours: Difficult to match, common across macro models.
  - Possibly, missmatch between theoretical concept and empirical counterpart due to heterogeneities on the labor market (Solon et al 1994); call for compositionally-adjusted measure of hours worked (Basu and Bundick 2017)
- Uncertainty persistence: unchanged

#### DSGE model: Assessment (cont'd)

► Crucial parameter: Higher risk aversion in Great Recession, which implies (Swanson 2018 ► formula):

$$RRA_{GR} = 145 > RRA_{lin} = 105$$

- Higher RRA in Great Recession in line with:
  - macro-finance (Cochrane 2017, Kim 2014)
  - experimental evidence (Cohn et al. 2015, Guiso et al. 2017)
  - surveys (Schildberg-Hörisch 2018)
- ► Nonlinearity implicit in the model not sufficient to match empirical facts ( no endog asymmetries )
- ▶ [Too high RRA?
  - ► Rudebusch and Swanson (2012): estimated RRA of 110
  - Barillas, Hansen and Sargent (2009): high RRA / true model vs. low RRA / doubts on model

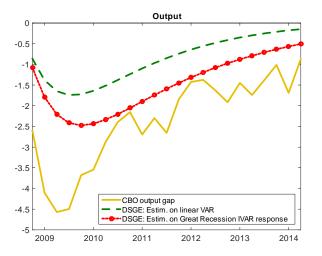


#### Role of nonlinearities for exercises with DSGE model

- ► How much of the total output loss during the Great Recession is due to an uncertainty shock?
- Uncertainty shock in 2008Q4: 4.4 standard deviation uncertainty shock (similar for linear/nonlinear VAR model)
- ▶ DSGE model's predictions conditional on estimates obtained with linear vs. nonlinear VAR?
- ► Actual data-reference: CBO output gap (2008Q3=0), sample: 2008Q4-2014Q2

#### Great recession: Role of nonlinearities

Output response to the uncertainty shock in 2008Q4



► Total loss: -53%; DSGE/nonlinear (linear) VAR: -31% (-20%)

#### Our findings offer empirical support to:

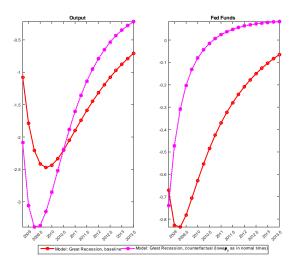
- "[...] Although some nonlinearities are accounted for in our modeling exercises, we cannot be certain that our simulations provide reasonable approximations of the economy's behavior in times of large idiosyncratic shocks."
  - ► Alan Greenspan, opening remarks at the "Monetary Policy and Uncertainty" symposium, Jackson Hole, August 29, 2003
- "For negative economic shocks such as the financial crisis in 2008–2009, the evidence rather consistently documents an increase in risk aversion, using a variety of methods."
  - Schildberg-Hörisch (2018, Journal of Economic Perspectives, "Are Risk Preferences Stable?")

## Output loss: Policy counterfactual

#### Role of monetary policy according to the DSGE

- What was the role of monetary policy in the propagation of the 2008Q4 uncertainty shock?
- Estimation:
  - $\rho_{\pi}^{GR} = \rho_{\pi}^{linear}$   $\rho_{y}^{GR} = 0.28 > \rho_{y}^{linear} = 0.20$
- Q: Did the Fed helped mitigate the depth of the Great Recession?
- Parallel with Great Depression: monetary policy was not accommodative enough (Friedman and Schwartz, 1963; Christiano, Motto, and Rostagno, 2003)
- ▶ Counterfactual ex.: replace  $ho_y^{GR}$  with  $ho_y^{linear}$ 
  - what would have happened?
- Actual data-reference: in 2008Q3 FFR = 1.94

#### Great recession: Role of monetary policy



► The Fed immediate and aggresive response played a significant role in mitigating the depth of the Great Recession.

#### Conclusions

- ► IVAR analysis, real effects of financial uncertainty shocks stronger during the Great Recession
- Predictions matched with a nonlinear DSGE model estimated in a state-dependent fashion. Crucial parameter: risk aversion (counter-cyclical). Informative to build theoretical models with endogenous mechanisms
- ▶ DSGE model estimated conditional on nonlinear VAR explains about 60% of the output loss during the Great Recession, just 40% if data modeled with a linear VAR
- Implications for model calibration and policy should a central bank directly react to financial volatility?
- More aggressive monetary policy prevented a greater recession
- Future research on financial & labor market frictions.

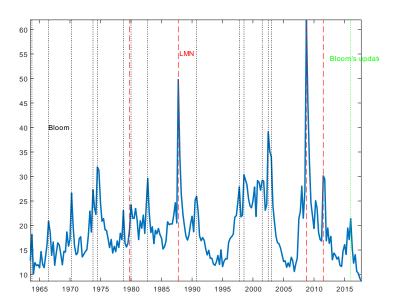
# Thanks a lot!

#### Uncertainty in the time of the Great Recession

"If you think that another Depression might be around the corner, better to be careful and save more. Better to wait and see how things turn out. Buying a new house, a new car or a new laptop can surely be delayed a few months. The same goes for firms: given the uncertainty, why build a new plant or introduce a new product now? Better to pause until the **smoke clears**. This is perfectly understandable behavior on the part of consumers and firms—but behaviour which has led to a collapse of demand, a collapse of output and the deep recession we are now in."

Olivier Blanchard, IMF's Chief Economist, The Economist, Jan. 29, 2009 • back

#### **Event constraints**

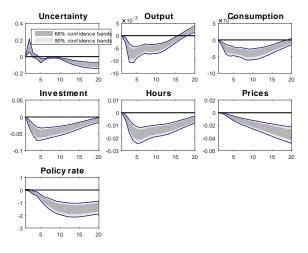


# Event constraints: cont'd ( back )

|        | Event contraints                           | 6   | Source        |
|--------|--|---|---------------|
| t      | Event                                      | Constraint on e <sub>FU.t</sub>                         |               |
| 1962Q4 | Cuban missile crisis                       | $e_{FU.t} > perc(\mathbf{e}_{FUt}(\mathbf{B}), 50th)$   | Bloom         |
| 1963Q4 | Assassination of JFK                       | $e_{FU.t} > perc(\mathbf{e}_{FUt}(\mathbf{B}), 50th)$   | Bloom         |
| 1966Q3 | Vietnam buildup                            | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 1970Q2 | Cambodia and Kent state                    | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 1973Q4 | OPEC I, Arab-Israeli War                   | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 1974Q3 | Franklin National                          | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 1978Q4 | OPEC II                                    | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 1979Q4 | Volcker experiment                         | $e_{FII,t} > perc(\mathbf{e}_{FIIt}(\mathbf{B}), 50th)$ | LMN           |
| 1980Q1 | Afghanistan, Iran hostages                 | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 1982Q4 | Monetary policy turning point              | $e_{FII,t} > perc(\mathbf{e}_{FIIt}(\mathbf{B}), 50th)$ | Bloom         |
| 1987Q4 | Black Monday                               | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 75th)$ | Bloom & LMN   |
| 1990Q4 | Gulf War I                                 | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 1991Q4 | Dissolution of the Soviet Union            | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 1997Q4 | Asian crisis                               | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 1998Q3 | Russian, LTCM default                      | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 2001Q3 | 9/11                                       | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 2002Q3 | Worlcom, Enron                             | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 2003Q1 | Iraq invasion                              | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom         |
| 2008Q4 | Great recession                            | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 75th)$ | Bloom & LMN   |
| 2011Q3 | Debt ceiling crisis                        | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | LMN           |
| 2016Q1 | FFR liftoff, China, Japanese negative rate | $e_{FIJ,t} > perc(\mathbf{e}_{FIJt}(\mathbf{B}), 50th)$ | Bloom (update |

| External variable constraints       |   |     |
|-------------------------------------|---|-----|
| External variable $S_t$             | Constraint on $\rho(e_{FUt}, S_t)$                      |     |
| Stock market return                 | $< perc(\rho(\mathbf{e}_{FUt}(\mathbf{B}), S_t), 50th)$ | LMN |
| Real price of gold (log difference) | $> perc(\rho(\mathbf{e}_{FUt}(\mathbf{B}), S_t), 50th)$ | LMN |

# Difference of state-conditional GIRFs: Model uncertainty

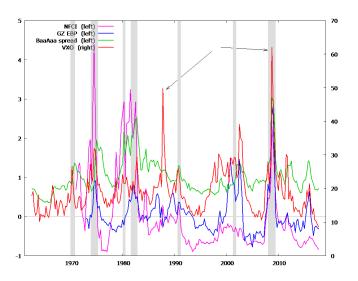




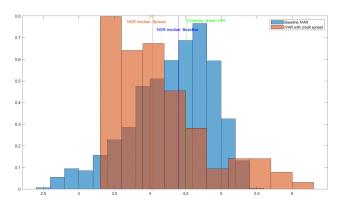
#### VAR checks: credit spread

- "the shocks that produced the [great] recession were primarily associated with financial disruptions and heightened uncertainty" (Stock and Watson, 2012)
- Narrative restrictions to disentangle the shocks:
  - uncertainty shocks: baseline NSR
  - financial (first moment) shocks: shock<sub>2008Q4</sub> > median(shocks(B)<sub>2008Q4</sub>) & shock<sub>1987Q4</sub> < median(shocks(B)<sub>1987Q4</sub>)
- NSR allow us to not take a stance on the contemporaneous relationship between uncertainty and financial shocks
- ▶ We are interested in correctly identifying uncertainty shocks

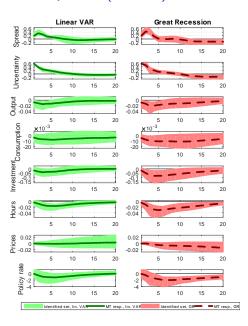
▶ Uncertainty vs financial stress:

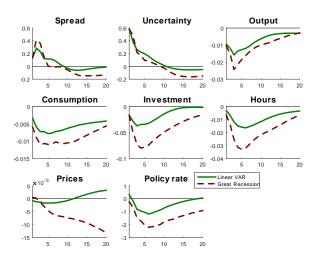


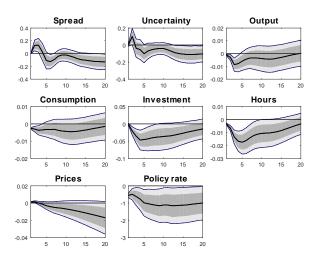
► Size of the uncertainty shock in 2008Q4:



▶ Part of the uncertainty spike is acknowledged to be a consequence of other shocks ( ► back\*)







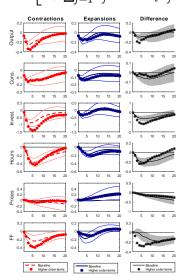
# Peak responses: Baseline vs. credit spread case

|                                    | Output | Consumption | Investment | Hours  |
|------------------------------------|--------|-------------|------------|--------|
| Baseline IVAR and VAR              |        |             |            |        |
| Peak response: Linear              | -1.54% | -1.03%      | -3.34%     | -1.90% |
| Peak response: Great Recession     | -2.32% | -1.36%      | -8.32%     | -3.75% |
| Ratio GR/Linear                    | 1.50   | 1.32        | 2.49       | 1.97   |
| IVAR and VAR with financial spread |        |             |            |        |
| Peak response: Linear              | -1.58% | -0.78%      | -3.70%     | -1.63% |
| Peak response: Great Recession     | -2.44% | -1.10%      | -8.11%     | -3.33% |
| Ratio GR/Linear                    | 1.54   | 1.41        | 2.19       | 2.04   |

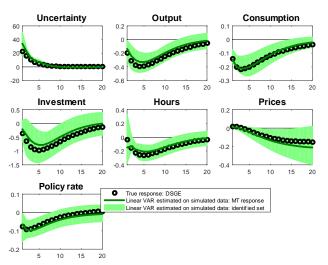


## IVAR with higher order terms Plack

$$Y_{t} = \alpha + \sum_{j=1}^{L} A_{j} Y_{t-j} + \begin{bmatrix} \sum_{j=1}^{L} c_{j} \ln VXO_{t-j} \times \Delta \ln GDP_{t-j} \\ + \sum_{j=1}^{L} c_{j} (\ln VXO_{t-j})^{2} \times \Delta \ln GDP_{t-j} \\ + \sum_{j=1}^{L} c_{j} \ln VXO_{t-j} \times (\Delta \ln GDP_{t-j})^{2} \end{bmatrix} + u_{t}$$



# Monte Carlo exercise: Validation identification in population

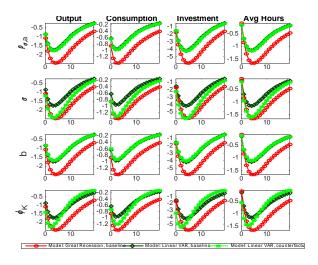


### DSGE calibrated parameters

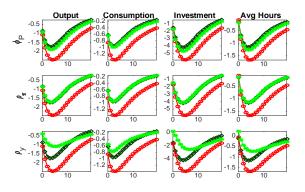
| Par.                | Description                                     | Value                  | Source         |
|---------------------|---|------------------------|----------------|
| $\sigma_{\sigma^a}$ | volatility of the uncertainty shock             | 0.004                  | BB (2018)      |
| $ ho^{a}$           | persistence of the preference shock             | 0.98                   | BB (2018)      |
| $\sigma^{a}$        | volatility of the preference shock              | 0.005                  | BB (2018)      |
| $\rho^Z$            | persistence of the technology shock             | 0.35                   | BB (2018)      |
| $\sigma^Z$          | volatility of the technology shock              | 0.019                  | BB (2018)      |
| α                   | capital's share in production                   | 0.333                  | BB (2017,2018) |
| β                   | household discount factor                       | 0.994                  | BB (2017,2018) |
| δ                   | steady state depreciation rate                  | 0.025                  | BB (2017,2018) |
| $\delta_1$          | first-order utilization parameter               | $1/\beta - 1 + \delta$ | BB (2017,2018) |
| П                   | steady state inflation rate                     | 1.005                  | BB (2017,2018) |
| ν                   | firm leverage parameter                         | 0.9                    | BB (2017,2018) |
| $\delta_2$          | second-order utilization parameter              | 0.0003                 | BB (2017,2018) |
| $\theta_{\mu}$      | elasticity of subst. between intermediate goods | 6.0                    | BB (2017,2028) |
| $\dot{\psi}$        | intertemporal elasticity of substitution        | 0.5                    | BB (2018)      |
|                     |   |                        |                |



# Counterfactual exercise: drivers response in great recession



# Counterfactual exercise: drivers response in great recession (2)







# Avg. distance criterion: drivers of state-dependent resp.

$$\textit{Avg.Crit} = \sum_{\textit{i} = \textit{contr}, \texttt{exp}} \frac{1}{2} \left( \widehat{\boldsymbol{\psi}^{\textit{i}}} - \boldsymbol{\psi} \left( \boldsymbol{\zeta}^{\textit{i}} \right) \right)^{'} \boldsymbol{\mathsf{V}}_{\textit{i}}^{-1} \left( \widehat{\boldsymbol{\psi}^{\textit{i}}} - \boldsymbol{\psi} \left( \boldsymbol{\zeta}^{\textit{i}} \right) \right)$$

| Average distance criterion:  |        |  |
|------------------------------|--------|--|
| 1 parameter-only estimations |        |  |
| $\rho_{\sigma^a}$            | 115.34 |  |
| $\sigma$                     | 108.06 |  |
| b                            | 113.67 |  |
| $\phi_K$                     | 110.13 |  |
| $\phi_P$                     | 118.40 |  |
| $ ho_{\pi}$                  | 121.82 |  |
| $\rho_{\rm v}$               | 114.95 |  |
|                              |        |  |



#### RRA for the Basu and Bundick's model with ext. habits

$$RRA \equiv_{\bar{N} \text{ arbitrary}} R^c$$

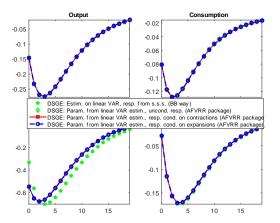
$$\begin{array}{ll} \textit{RRA} & = & \left(\frac{\eta}{\eta + \left(1 - \eta\right)\left(1 - b\right)}\right) \cdot \\ & \left(\frac{1}{\psi} \cdot \frac{\left(1 + \frac{\left(1 - \eta\right)}{\eta}\left(1 - b\right)\right)}{\left(1 - b\right)\left(1 + \frac{\left(1 - \eta\right)}{\eta}\right)} + \left(\sigma - \frac{1}{\psi}\right)\left(\frac{\eta}{\left(1 - b\right)} + 1 - \eta\right)\right) \end{array}$$

• if 
$$b = 0 \rightarrow RRA = \eta \cdot \sigma$$



#### Any endog. asymmetry?: DSGE state-conditional GIRFs

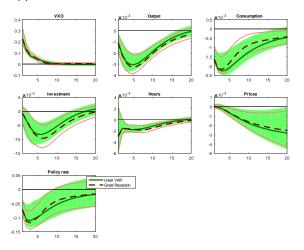
Role of initial conditions in the pruned DSGE model for the propagation of the uncertainty shock (Andreasen, Fernández-Villaverde, and Rubio-Ramírez (2017)):



Theoretically proved in Cacciatore and Ravenna (2018)

# Any endog. asymmetry?: (I)VAR estim. on simulated data

Sample of 2500 simulated observations from non-pruned, III order-approx. DSGE:



► If non-pruned, IV order-approx. DSGE: approximation errors too big! ( ► back )